

English

## SOLUTION SKETCH FOR

# TMA4285 TIME SERIES MODELS 11. December 2008 Time: 09:00–13:00

## Problem 1

- a) The ACF  $\rho(k) = \rho_k$  for the given AR(1) process is given as  $\rho_k = (-\phi)^k$  for k = 0, 1, 2, ...From the figures one may therefore read off the values of  $\phi$ : 1)  $\phi = -0.9$ , 2)  $\phi = -0.4$ , 3)  $\phi = 0.8$ , 4)  $\phi = 0.5$
- b) The plot of  $Y_t$  versus  $Y_{t-1}$  displays a significant negative correlation, while the plot of  $Y_t$  versus  $Y_{t-2}$  indicates a weak positive correlation (draw a vertical and horizontal line through the origin and study each of the quadrants). No such trend can be detected in the plot of  $Y_t$  versus  $Y_{t-3}$ . This indicates that the data is generated by an MA(2) model.
- c) The recorded time series displays both increasing variability (variance), trend and seasonality. Therefore, in order to stabilize the variability, a transformation of the data is suggested. This could be done by the Cox-Box transformation, a log transformation is a typical choice. After an appropriate transformation, the trend is removed by differencing. Here one differencing would seem sufficient. Finally, the seasonal component may be removed by s-differencing, where s denotes the identified seasonal period. The residual process after these operations can then be studied for a possible ARMA model fit.

## Problem 2

a) For  $X_t$  to be an ARMA(2,1) process it has to be stationary, and the AR and MA polynomials cannot have common roots. Stationarity is guaranteed by  $\phi(z) \neq 0$  for |z| = 1 ( $z \in \mathbf{C}$  = the complex numbers). That is, the roots of the AR polynomial  $\phi(z) = 1 - z + \phi^2 z^2$  cannot lie on the unit circle. For  $\phi = 0$ , obviously  $\phi(z) = 0$  for z = 1.

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Hence, a necessary condition for  $X_t$  to be stationary is that  $\phi \neq 0$ . For  $\phi \neq 0$ , i.e.  $\phi^2 > 0$ , the roots of  $\phi(z)$  are obtained as,

$$z^{\pm} = \frac{1 \mp \sqrt{1 - 4\phi^2}}{2\phi^2} = \frac{2}{1 \pm \sqrt{1 - 4\phi^2}}$$

Can  $|z^{\pm}| = 1$  for some value of  $\phi \neq 0$ ? It is seen that  $z^{\pm}$  are real for  $0 < \phi^2 \leq 1/4$ , and that  $|z^{\pm}| > 1$ . For  $\phi^2 > 1/4$ , there will be two complex conjugate roots  $z^{\pm} = 2/(1\pm i\sqrt{4\phi^2 - 1})$  ( $i = \sqrt{-1}$ ). It follows that  $|z^{\pm}| = 1$  if  $4 = 1 + (4\phi^2 - 1) = 4\phi^2$ , which is satisfied for  $\phi = \pm 1$ . Conclusion:  $X_t$  is stationary for  $\phi \notin \{0, \pm 1\}$ 

The requirement of no common roots need only be checked when  $0 < \phi^2 \le 1/4$ . The root of the MA polynomial is  $z = -1/\theta$ . Hence, there will be no common roots when

$$\theta \neq \frac{-1 \pm \sqrt{1 - 4\phi^2}}{2}, \text{ for } 0 < \phi^2 \le 1/4,$$

else  $\theta$  is arbitrary.

b) For  $X_t$  to be causal,  $|z^{\pm}| > 1$ . According to a) we only need to investigate the parameter range  $\phi^2 > 1/4$ . In this range, it follows that  $|z^{\pm}| > 1$  if and only if  $\phi^2 < 1$ . Hence,  $X_t$  is causal if and only if  $0 < \phi^2 < 1$ .

 $X_t$  is invertible if and only if  $1 + \theta z \neq 0$  for  $|z| \leq 1$ , that is, for  $|1/\theta| > 1$ . Hence,  $X_t$  is invertible if and only if  $|\theta| < 1$ .

- c) For the AR(2) process at hand there will be three cases to consider:
  - 1. For  $0 < \phi^2 < 1/4$ , and  $h \in \mathbf{Z}$ ,

$$\gamma(h) = c_1 \left(\frac{1+\sqrt{1-4\phi^2}}{2}\right)^{|h|} + c_2 \left(\frac{1-\sqrt{1-4\phi^2}}{2}\right)^{|h|},$$

where  $c_1$  and  $c_2$  are two real constants.

2. For  $\phi^2 = 1/4$ , and  $h \in \mathbf{Z}$ ,

$$\gamma(h) = (c_1 + c_2|h|)2^{-|h|}$$

where  $c_1$  and  $c_2$  are two real constants.

3. For  $1/4 < \phi^2 < 1$ , and  $h \in \mathbf{Z}$ ,

$$\gamma(h) = c \left(\frac{1 + i\sqrt{4\phi^2 - 1}}{2}\right)^{|h|} + \overline{c} \left(\frac{1 - i\sqrt{4\phi^2 - 1}}{2}\right)^{|h|},$$

where  $c = c_1$  is a complex number in general, and  $c_2 = \overline{c}$ , which is the complex conjugate of c. This is necessary to make  $\gamma(h)$  real. Using polar representation of a complex number, we may write  $c = (a/2) e^{ib}$  for suitable real numbers a and b. Also

$$\frac{1 + i\sqrt{4\phi^2 - 1}}{2} = |\phi| e^{i\theta} ,$$

where  $\tan \theta = \sqrt{4\phi^2 - 1}$ . It is then obtained that

$$\gamma(h) = a|\phi|^{|h|}\cos(\theta|h| + b),$$

d) Using the relation  $X_t = X_{t-1} - \frac{1}{2}X_{t-2} + Z_t$ , and calculating  $E[X_tX_t]$  and  $E[X_tX_{t+1}]$ , the following two equations are obtained,

$$\gamma(1) = \frac{1}{4}\gamma(0) + \sigma^2$$

and

$$\gamma(2) = \frac{1}{2}\gamma(1) - \gamma(0) + 2\sigma^2 = -\frac{7}{8}\gamma(0) + \frac{5}{2}\sigma^2$$

From the previous point we know that

$$\gamma(h) = c \left(\frac{1+\mathrm{i}}{2}\right)^{|h|} + \overline{c} \left(\frac{1-\mathrm{i}}{2}\right)^{|h|},$$

which leads to the equations

$$c \frac{1+i}{2} + \overline{c} \frac{1-i}{2} = \frac{1}{4}(c+\overline{c}) + \sigma^2,$$

and

$$c\mathbf{i} - \overline{c}\mathbf{i} = -\frac{7}{8}(c+\overline{c}) + \frac{5}{2}\sigma^2.$$

Solving the equations gives the solution

$$c = \frac{4}{13}(7+2i)\sigma^2 = \frac{4}{13}\sqrt{53}e^{i0.278},$$

where  $0.278 = \tan^{-1}(2/7)$ . Noting that  $(1 + i)/2 = (1/\sqrt{2}) e^{i\pi/4}$ , it follows that

$$\gamma(h) = \frac{8}{13}\sqrt{53}\,\sigma^2 \left(\frac{1}{2}\right)^{|h|/2} \cos(\frac{\pi}{4}|h| + 0.278)\,,$$

The ACF then becomes

$$\rho(h) = \frac{\cos\left(\frac{\pi}{4}|h| + 0.278\right)}{2^{|h|/2}\cos(0.278)}, \ h \in \mathbf{Z}$$
(1)

#### Problem 3

a) Since  $\varepsilon_t$  is IID, it follows that  $\varepsilon_t$  is independent of  $\varepsilon_s$  for every s < t. This implies that  $\varepsilon_t$  is independent of  $\varepsilon_{t-1}^2 \cdot \ldots \cdot \varepsilon_{t-j}^2$  for every  $j = 1, 2, \ldots$ . Except for constants, H(t) consists of a sum of such terms, and therefore  $\varepsilon_t$  and  $H_t$  are independent random variables for each t. Then  $\varepsilon_t$  and  $\sqrt{H_t}$  are also independent random variables for each t. It is then obtained that,  $E[X_t] = E[\varepsilon_t] E[\sqrt{H_t}] = 0$ ,

$$\mathbf{E}[X_t^2] = \mathbf{E}[\varepsilon_t^2 H_t] = \mathbf{E}[\varepsilon_t^2] \mathbf{E}[H_t] = \mathbf{E}[H_t]$$
$$= \alpha_0 \left(1 + \sum_{j=1}^{\infty} \alpha_1^j \mathbf{E}[\varepsilon_{t-1}^2] \cdot \ldots \cdot \mathbf{E}[\varepsilon_{t-j}^2]\right) = \alpha_0 \left(1 + \sum_{j=1}^{\infty} \alpha_1^j\right) = \frac{\alpha_0}{1 - \alpha_1}.$$

Since  $\varepsilon_t$  is IID, it is seen that  $X_t$  is strictly stationary. Hence, moments, when they exist, are independent of t. From the equation  $X_t^2 = \varepsilon_t^2(\alpha_0 + \alpha_1 X_{t-1}^2)$ , it follows that

$$X_t^4 = \varepsilon_t^4 \left( \alpha_0^2 + 2\alpha_0 \alpha_1 X_{t-1}^2 + \alpha_1^2 X_{t-1}^4 \right),$$

Hence, if  $m_4 = \mathbb{E}[X_t^4]$  exists, it must satisfy the equation  $(\mathbb{E}[\varepsilon_t^4] = 3)$ ,

$$m_4 = 3\left(\alpha_0^2 + 2\alpha_0\alpha_1\frac{\alpha_0}{1-\alpha_1} + \alpha_1^2m_4\right),\,$$

This leads to the equation

$$m_4 = \frac{3\,\alpha_0^2(1+\alpha_1)}{(1-\alpha_1)(1-3\alpha_1^2)}\,.$$

This equation can only be satisfied if  $0 < \alpha_1^2 < 1/3$ , which becomes the condition for finite  $m_4$ .

b) By a similar argument as above, it is seen that  $\eta_t$  is a strictly stationary process. By the assumption that  $E[X_t^4] < \infty$ , it follows that  $\eta_t$  has finite second order moments. It is therefore a (weakly) stationary process.

$$\mathbf{E}[\eta_t] = \mathbf{E}[X_t^2] - \mathbf{E}[H_t] = 0.$$

and for  $h \ge 1$ ,

$$E[\eta_{t+h} \eta_t] = E[(\varepsilon_{t+h}^2 - 1)(\varepsilon_t^2 - 1) H_{t+h} H_t] = E[(\varepsilon_{t+h}^2 - 1)] E[(\varepsilon_t^2 - 1) H_{t+h} H_t] = 0.$$

since  $\varepsilon_{t+h}$  is independent of  $\varepsilon_s$  for every s < t+h, and therefore  $\varepsilon_{t+h}^2 - 1$  is independent of  $(\varepsilon_t^2 - 1) H_{t+h} H_t$ . It follows that  $\eta_t$  is white noise with variance

$$\sigma_0^2 = \mathbf{E}[\eta_t^2] = \mathbf{E}[(\varepsilon_t^2 - 1)^2 H_t^2] = \mathbf{E}[(\varepsilon_t^2 - 1)^2] \mathbf{E}[H_t^2] = \mathbf{E}[\varepsilon_t^4 - 2\varepsilon_t^2 + 1] \frac{\mathbf{E}[X_t^4]}{\mathbf{E}[\varepsilon_t^4]} = \frac{2}{3}m_4 = \frac{2\alpha_0^2(1 + \alpha_1)}{(1 - \alpha_1)(1 - 3\alpha_1^2)}$$

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c) From the definition of  $\eta_t$ , it follows that

$$X_t^2 = H_t + \eta_t = \alpha_0 + \alpha_1 X_{t-1}^2 + \eta_t.$$
 (2)

Introducing the process  $Y_t = X_t^2 - \alpha_0/(1 - \alpha_1)$ , it is obtained that,

$$Y_t = \alpha_1 Y_{t-1} + \eta_t. \tag{3}$$

Since  $0 < \alpha_1 < 1/\sqrt{3}$ , it is seen that  $\phi(z) = 1 - \alpha_1 z \neq 0$  for  $|z| \le 1$ . It follows that  $Y_t$  becomes a causal AR(1) process, and therefore also  $X_t^2$  (in the non-zero mean form).